# Big-Data Algorithms: 

Computing the $\ell_{2}$ Norm
Reference: http://www.sketchingbigdata.org/fall17/lec/lec3.pdf

## Basic Data Stream Model

- Single pass over the data $i_{1}, i_{2}, \ldots, i_{n} \in[m]$. Here, $[m]=\{1,2, \ldots, m\}$.


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- Randomness and approximation OK (almost always necessary)
- Last lecture: estimating the number of distinct elements

$$
\mathbb{P}(\text { relative error }<\varepsilon)>\frac{2}{3}
$$

with space $O\left(\log n+(1 / \varepsilon)^{2}\right)$.

## Generalization

- Vector $x=\left(x_{1}, x_{2}, \ldots, x_{m}\right)$, where each $x_{i}$ is number of times $i \in[m]$ has been seen so far.


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- Johnson-Lindenstrauss


## Why Calculate $\ell_{2}$ norm?

- General $\ell_{p}$ norms:

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\|x\|_{\ell_{p}}= \begin{cases}\left(\sum_{i=1}^{n}\left|x_{i}\right|^{p}\right)^{1 / p} & \text { if } p<\infty \\ \max _{1 \leq i \leq n}\left|x_{i}\right| & \text { if } p=\infty\end{cases}
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- $\|x\|_{\ell_{2}}$ measures spikiness of $x$ :

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- $\ell_{2}$-norm estimation also arises in database applications.


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- Choose $r_{1}, \ldots, r_{m}$ to be independently identically distributed random variables, with

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Z=\langle r, x\rangle=\sum_{i=1}^{n} r_{i} x_{i}
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under increments to the $x_{i}$.
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\text { update } x \text { by }(i, a) \text { simply via } Z \leftarrow Z+r_{i} a
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- Analysis?
- Compute expectation of $Z^{2}$.
- Bound the variance of $Z^{2}$.


## AMS Algorithm: Compute Expectation

- We have

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\mathbb{E}\left[Z^{2}\right] & =\mathbb{E}\left(\sum_{i=1}^{n} r_{i} x_{i}\right)^{2}=\mathbb{E}\left(\sum_{i, j=1}^{n} r_{i} x_{i} r_{j} x_{j}\right) \\
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- So

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\mathbb{E}\left[Z^{2}\right]=\sum_{i=1}^{n} x_{i}^{2}=\|x\|_{\ell_{2}}^{2}
$$

and hence $Z^{2}$ is an unbiased estimator of $\|x\|_{\ell_{2}}^{2}$.

## AMS Algorithm: Bounding the Variance

- We have

$$
\operatorname{Var}\left[Z^{2}\right]=\mathbb{E}\left[\left(Z^{2}\right)^{2}\right]-\left(\mathbb{E}\left[Z^{2}\right]\right)^{2}=\mathbb{E}\left[Z^{4}\right]-\|x\|_{\ell_{2}}^{4}
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- $6 \sum_{i, j=1}^{n}\left(r_{i} r_{j} x_{i} x_{j}\right)^{2}$, with expectation $6 \sum_{1 \leq i<j \leq n} x_{i}^{2} x_{j}^{2}$


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- Terms involving no repeated multipliers (such as $\left.r_{1} x_{1} r_{2} x_{2} r_{3} x_{3} r_{4} x_{4}\right)$, with expectation 0 .
- So

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\mathbb{E}\left[Z^{4}\right]=\sum_{i=1}^{n} x_{i}^{4}+6 \sum_{1 \leq i<j \leq n}^{n} x_{i}^{2} x_{j}^{2}
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## AMS Algorithm: Bounding the Variance (con'td)

- Recall that

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& =\sum_{i=1}^{n} x_{i}^{4}+6 \sum_{1 \leq i<j \leq n}^{n} x_{i}^{2} x_{j}^{2}-\left(\sum_{i=1}^{n} x_{i}^{4}+2 \sum_{1 \leq i<j \leq n}^{n} x_{i}^{2} x_{j}^{2}\right) \\
& =4 \sum_{1 \leq i<j \leq n}^{n} x_{i}^{2} x_{j}^{2} \leq 2\left(\sum_{i=1}^{n} x_{i}^{2}\right)^{2}=2\|x\|_{\ell_{2}}^{2}
\end{aligned}
$$

## AMS Algorithm: Completing the Analysis

- We have an estimator $Z^{2} \approx\|x\|_{\ell_{2}}^{2}$, with $E\left[Z^{2}\right]=\|x\|_{\ell_{2}}^{2}$ and $\sigma=\operatorname{Var}\left[Z^{2}\right] \leq 2\|x\|_{\ell_{2}}^{2}$.


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- Apply Chebyshev inequality

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\mathbb{P}[|\mathbb{E}[Y]-Y| \geq c \sigma] \leq \frac{1}{c^{2}} \quad \forall c>0
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to find

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\mathbb{P}\left[\left|\mathbb{E}\left[Z^{2}\right]-\|x\|_{\ell_{2}}^{2}\right| \geq c \sqrt{2}\|x\|_{\ell_{2}}^{2}\right] \leq \frac{1}{c^{2}}
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- Problem: This gives a lousy estimator if failure probability $\delta$ is small. For instance, if $\delta=\frac{1}{3}$, must choose $c=3$, finding

$$
\mathbb{P}\left[\left|\mathbb{E}\left[Z^{2}\right]-\|x\|_{\ell_{2}}^{2}\right| \geq 3 \sqrt{2}\|x\|_{\ell_{2}}^{2}\right] \leq \frac{1}{9}
$$

But $\mathbb{E}\left[Z^{2}\right] \geq 0$, so this is worse than natural bound.

## AMS+ Algorithm

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- Moreover,

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\operatorname{Var}[Y]=\frac{1}{k^{2}} \sum_{j=1}^{k} \operatorname{Var}\left[Z_{j}^{2}\right] \leq \frac{2}{k}\|x\|_{\ell_{2}}^{2},
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so that Chebyshev's inequality yields

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\mathbb{P}\left[\left|\mathbb{E}[Y]-\|x\|_{\ell_{2}}^{2}\right| \leq c \sqrt{2 / k}\|x\|_{\ell_{2}}^{2}\right]<1 / c^{2}
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- Space usage: $O\left(\log (m n) / \varepsilon^{2}\right)$ bits (not counting the $r_{i}$ )


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- Error bound too loose to be useful for small $\delta$ : For $c=O(1 / \sqrt{\delta})$, need $k=O\left(1 /\left(\delta \varepsilon^{2}\right)\right.$, linear in $\delta$. That's because we only used second moment.


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- Set $k=O\left(1 / \varepsilon^{2} \log (1 / \delta)\right)$ to get $1 \pm \varepsilon$ approximation with probability $1-\delta$.


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- Fast JL, sparse JL: reduce updating time

